

Bayesian regression models in gretl

L. Pedini

Dipartimento di Scienze Economiche e Sociali (DiSES)
Univpm - Facoltà di Economia G.Fuà

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The package BayTool

Baytool introduces common Bayesian model estimations in `gretl`.

At the moment, the package provides:

- homoskedastic linear model;
- heteroskedastic linear model;
- Bayesian LASSO;
- Latent variable models (under development).

The package includes:

- easy and immediate syntax (script) and a GUI;
- flexibility in priors;
- parallelization via MPI;
- plots and diagnostics.

The idea behind and the final aim

Providing a Bayesian alternative to each frequentist estimation command (native or not)

Why going Bayesian?

Pros:

- “more elegant and rigorous” (Koop, 2003);
- parameters are **random variables**;
- priors add information and regularize
- exploit Monte Carlo integration techniques.

Cons

- computationally more complex;
- sensitivity to prior choices;
- prior hyperparameters

Posterior

Most of the Bayesian interest lies in the posterior of a given parameter θ , given data \mathcal{D} :

$$P(\theta|\mathcal{D}) = \frac{p(\mathcal{D}|\theta)P(\theta)}{P(\mathcal{D})} \propto p(\mathcal{D}|\theta)P(\theta)$$

- $P(\theta)$, *prior*;
- $p(\mathcal{D}|\theta)$, *likelihood*;
- $P(\theta|\mathcal{D})$, **posterior**

An example: the linear model

$$y = X\beta + \varepsilon, \quad \varepsilon \sim N(0, \sigma^2 I_n)$$

$n \times 1$ $n \times k$ $n \times 1$

The likelihood is given by,

$$p(y|\beta, \sigma^2) = \left(\frac{1}{2\pi\sigma^2}\right)^{n/2} \exp\left[-\frac{1}{2\sigma^2}(y - X\beta)^T(y - X\beta)\right]$$

Priors

Rewrite the likelihood as,

$$p(y|\beta, \sigma^2) = \left(\frac{1}{2\pi}\right)^{n/2} \left\{ \frac{1}{\sigma} \exp\left[-\frac{1}{2\sigma^2}(\beta - \hat{\beta})^T(X^T X)(\beta - \hat{\beta})\right] \right\} \\ \times \left\{ \left(\frac{1}{\sigma^2}\right)^{\hat{a}/2} \exp\left[-\frac{\hat{a}s^2}{2\sigma^2}\right] \right\}$$

where:

$$\hat{\beta} = (X^T X)^{-1} X^T y \\ \hat{s}^2 = \frac{(y - X\hat{\beta})^T (y - X\hat{\beta})}{\hat{a}}, \quad \hat{a} = n - k$$

An obvious choice for the priors should follow a similar pattern:
 $\beta|\sigma^2 \sim \text{MVN}; \sigma^2 \sim \text{IG}$

Conjugate

$$P(\beta, \sigma^2) = P(\beta|\sigma^2)P(\sigma^2)$$

↓

$$\beta, \sigma^2 | y \sim \text{NIG}$$

$$\beta | y \sim \text{MVT}$$

$$\sigma^2 | y \sim \text{IG}$$

Independent

$$P(\beta, \sigma^2) = P(\beta)P(\sigma^2)$$

↓

$$\beta, \sigma^2 | y \sim ???$$

↓

Gibbs sampler

$$\beta | \sigma^2, y \sim \text{MVN}$$

$$\sigma^2 | \beta, y \sim \text{IG}$$

House price example (Koop, 2003)

Anglin and Gencay (1996) dataset on $n = 546$ house sold in Canada in 1987

- y , house sales price
- x_1 , lot size
- x_2 , number of bedrooms
- x_3 , number of bathrooms
- x_4 , number of storeys

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-----  
bundle ret = BT_LinMod(series y, list X, bundle prior, bundle opt)  
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```

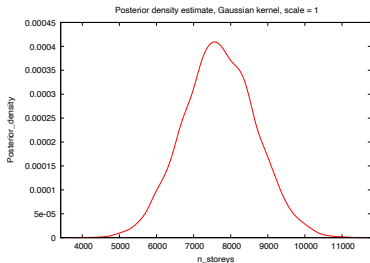
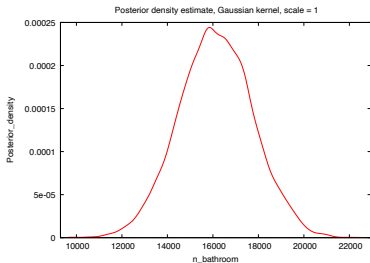
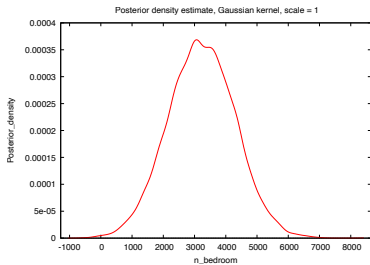
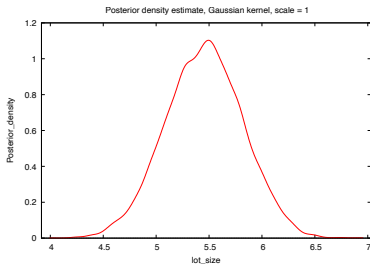
Dependent variable: y

Prior specification: Independent NIG prior

Estimation method: Gibbs Sampler - 10000 iter (1000 burn-in)

Summary statistics:

	Prior_m	Prior_se	NI-Post_m	NI-Post_se	I-Post_m	I-Post_se
const	0.00	10000.00	-4009.55	3609.79	-4131.25	3269.36
x1	10.00	5.00	5.42	0.37	5.45	0.37
x2	5000.00	2500.00	2824.61	1217.06	3224.68	1068.05
x3	10000.00	5000.00	17105.17	1737.65	16123.06	1624.84
x4	10000.00	5000.00	7634.90	1009.84	7691.85	974.03
Sig2	4e+07	6e+07	3e+08	2e+07	3e+08	2e+07



Shrinkage regression

Suppose to have a sample of $n = 300$ observations ($n_0 = 200$ train and $n_1 = 100$ test) of a given y and X , with $k = 300$.

$$y \quad n \times 1 \quad X \quad n \times k$$

$$DGP : y = Z\beta + \varepsilon, \quad Z = \{x_1, \dots, x_{75}\}$$

Under these circumstances common OLS **fails** when using X

- ℓ_1 norm - LASSO (Tibshirani, 1996)
- ℓ_2 norm - ridge

Ridge regression

Frequentist framework:

$$\min_{\beta} \left(\frac{1}{n_0} \|y - X\beta\|_2^2 + \lambda \sum_{j=1}^k \beta_j^2 \right) \rightarrow \hat{\beta}_r = (X^T X + \lambda I)^{-1} X^T y$$

Bayesian framework: introducing regularization is quite natural; priors deal the job.

$$\beta | \sigma^2 \sim \text{MVN}(0, \sigma^2 \gamma I) \rightarrow E(\beta | y) = \left(X^T X + \frac{1}{\gamma} I \right)^{-1} X^T y$$

LASSO

Frequentist framework (Tibshirani, 1996):

$$\min_{\beta} \left(\frac{1}{n_0} \|y - X\beta\|_2^2 + \lambda \sum_{j=1}^k |\beta_j| \right)$$

Bayesian framework (Park and Casella, 2008):

$$\beta | \sigma^2, \tau_1^2, \dots, \tau_k^2 \sim \text{MVN}(0, \sigma^2 D), \quad D = \text{diag}\{\tau_1^2, \dots, \tau_k^2\}$$

$$P(\tau_1^2, \dots, \tau_k^2) = \prod_{j=1}^k \frac{\lambda^2}{2} \exp(-\lambda^2 \tau_j^2 / 2)$$

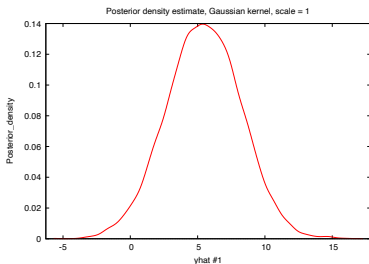
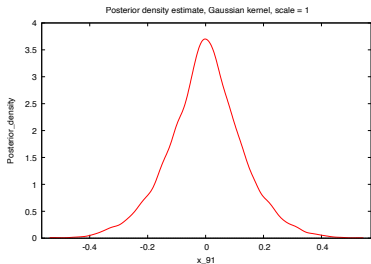
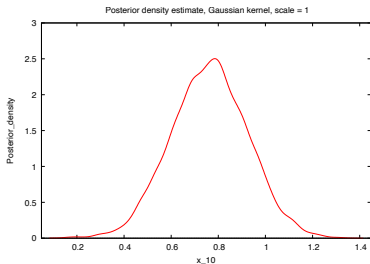
Comparing predictions on test set

	MSE(\hat{y})	CPU (sec)
ridge (cv)	10.45	0.42
Bay. ridge ($\gamma = 0.01$)	20.55	81.81
Bay. ridge ($\gamma = 0.05$)	11.17	79.19
Bay. ridge ($\gamma = 0.10$)	9.49	85.76
Bay. ridge ($\gamma = 1$)	9.50	78.86
Bay. ridge ($\gamma = 10$)	9.45	80.18
LASSO (cv)	3.37	5.56
Bay. LASSO (hyper)	4.80	77.74

Notes : cv = cross-validation with 10 folds, 50 λ sequences; hyper = hyperprior set-up. Posterior mean of sampled predictions used in Bayesian experiments with 10000 Gibbs iterations with 1000 burn-in (MPI with 4 threads).

Inference

In the Bayesian context deriving inference tools is straightforward.

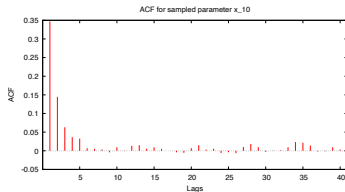
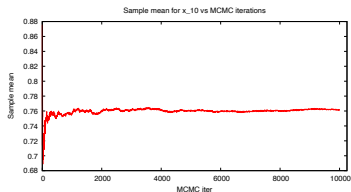
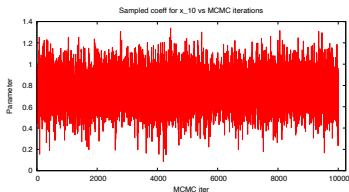


MCMC diagnostics

- autocorrelation/mixing “raw” indicators
 - time series plot, ACF;
 - effective sample size (Vats et al., 2019);
 - numerical standard errors.
- more “formal” testing
 - Heidelberger and Welch (1983);
 - Geweke (1992);
 - Gelman and Rubin (1992); Brooks and Gelman (1998).

Autocorrelation plots

Diagnostic plots from previous Bayesian LASSO example



Conclusion

- Baytool includes common Bayesian alternatives to frequentist commands;
- this becomes very useful when the frequentist alternatives are complex or lack from proper inferential tools;
- computational burden is dealt with MPI;
- Diagnostic tool available for MCMC cases.

References

- Anglin, P. M. and Gencay, R. (1996). Semiparametric estimation of a hedonic price function. *Journal of applied econometrics*, 11(6):633–648.
- Brooks, S. P. and Gelman, A. (1998). General methods for monitoring convergence of iterative simulations. *Journal of Computational and Graphical Statistics*, 7(4):434–455.
- Gelman, A. and Rubin, D. B. (1992). Inference from iterative simulation using multiple sequences. *Statistical Science*, pages 457–472.
- Geweke, J. (1992). Evaluating the accuracy of sampling-based approaches to the calculation of posterior moments. In *Bayesian Statistics (J.M. Bernardo, J.O. Berger, A.P. Dawid and A.F.M. Smith eds.)*, volume 4, pages 169–193. Clarendon Press, Oxford, UK.
- Heidelberger, P. and Welch, P. D. (1983). Simulation run length control in the presence of an initial transient. *Operations Research*, 31(6):1109–1144.
- Koop, G. (2003). *Bayesian Econometrics*. John Wiley & Sons Ltd.
- Park, T. and Casella, G. (2008). The bayesian lasso. *Journal of the American Statistical Association*, 103(482):681–686.
- Tibshirani, R. (1996). Regression shrinkage and selection via the lasso. *Journal of the Royal Statistical Society: Series B (Methodological)*, 58(1):267–288.
- Vats, D., Flegal, J. M., and Jones, G. L. (2019). Multivariate output analysis for markov chain monte carlo. *Biometrika*, 106(2):321–337.