# Bayesian Model Averaging

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- Publicly available software (some in gretl)
- Some recommendations and open questions

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- Model averaging: our inference is averaged over all the models in the model space considered, using weights that are either derived from Bayes' theorem (BMA) or from sampling-theoretic optimality considerations (FMA). Here focus on BMA

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E.g. wish to predict the unobserved  $y_f$  on the basis of the observed y. Sampling model for  $y_f$  and y jointly is  $p(y_f|y,\theta_j,M_j)p(y|\theta_j,M_j)$ , where  $M_j$  is the model selected from K possible models, and  $\theta_i \in \Theta_i$  are the parameters of  $M_i$ .

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Assign a (continuous) prior  $p(\theta_j|M_j)$  for the parameters and a discrete prior  $P(M_j)$  on the model space. Predictive distribution is

$$p(y_f|y) = \sum_{j=1}^K \left[ \int_{\Theta_j} p(y_f|y,\theta_j,M_j) p(\theta_j|y,M_j) d\theta_j \right] P(M_j|y) \quad (1)$$

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Averaging at two levels: over parameter values, given each possible model, and discrete averaging over all possible models

Square brackets in (1): predictive given  $M_j$  obtained using the posterior of  $\theta_j$  given  $M_j$ , which is

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with the second equality defining  $p(y|M_j)$ , used in computing the posterior probability of  $M_j$ :

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Denominators of both averaging operations are made explicit in (2) and (3).  $p(y|M_j)$  in (2) is the marginal likelihood of  $M_j$  and is a key quantity: Bayes factor is the ratio of marginal likelihoods (posterior odds = Bayes factor \* prior odds). p(y) in (3) is a sum (challenge often lies in the number of models K).

More generally, the posterior distribution of any quantity of interest, say  $\Delta$ , which has a common interpretation across models is a mixture of the model-specific posteriors with the posterior model probabilities as weights, *i.e.* 

$$P_{\Delta|y} = \sum_{j=1}^{K} P_{\Delta|y,M_j} P(M_j | y).$$
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Growing importance of model averaging as a solution to model uncertainty is illustrated by Figure 1, which plots the citation profile over time in the literature. The figure also indicates influential papers (with 250 citations or more) published in either economics or statistics journals

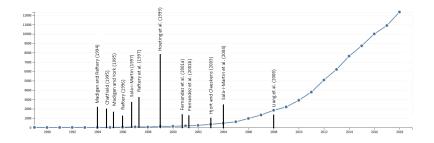


Figure: Total number of citations to papers with topic "model averaging" over years 1989-2018. Papers in economics or statistics journals with at least 250 citations are indicated by vertical lines proportional to the number of citations received. Source: Web of Science, Jan. 29, 2019.

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- Theory: which variables are important drivers? Often, theories regarding variable inclusion do not contradict each other ("open-endedness" of the theory)
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Usually theoretical results are derived under the assumption that  $\mathcal M$  contains the "true" data-generating model (" $\mathcal M$ -closed"), but most important results like model selection consistency extend to " $\mathcal M$ -open" settings in an intuitive manner.

Most common setting: model uncertainty about which covariates to include, *i.e.* under model j the n obs. in y are generated from

$$y|\theta_j, M_j \sim N(\alpha \iota + Z_j \beta_j, \sigma^2).$$
 (5)

Here  $\iota$  is a vector of ones,  $Z_j$  groups  $k_j$  of the possible k regressors and  $\beta_j \in \Re^{k_j}$  are the regression coefficients. All models contain an intercept  $\alpha \in \Re$  and a scale  $\sigma > 0$  with a common interpretation.

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Economics: k up to 100 (growth), so  $K > 10^{30}$  and need efficient computational tools. Genetics (usually n << k): k could be up to 100,000, leading to  $K > 10^{30,000}$ !

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Prior over models:

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Only requires choice of two scalars g and w: hyperpriors are recommended (adaptive and more robust)

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- BMA predicts at least as well as any single model (assuming data is generated by (1)) and there is ample empirical evidence for clear superiority (probabilistic forecasts)

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So no need to avoid using large  $\mathcal{M}!$ 

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Hyperpriors on g and w can have a large effect on the induced penalties for model complexity but not on the impact of the relative fit of the models

#### Complexity penalties

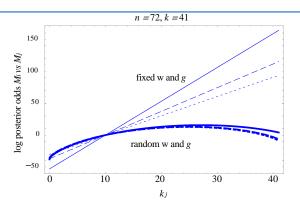


Figure: Posterior odds as a function of  $k_j$  when  $k_i=10$  with equal fit, using prior mean model size m=7 (solid), m=k/2 (dashed), and m=2k/3 (dotted). Bold lines correspond to random w and g

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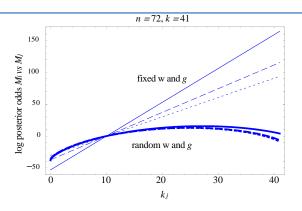


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Hyperpriors more robust to m, less extreme and penalize models of size around k/2 (multiplicity penalty)

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Hybrids of frequentist and Bayesian methods were used e.g. to deal with endogenous regressors: BIC approximations to posterior model probabilities for averaging over classical two-stage least squares (2SLS) estimates.

# Other sampling models

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#### BMA for many other models has been considered:

- Generalized linear models (GLMs), for example logistic, probit or ordered response models
- Generalized additive models (nonlinear effects)
- Models for outliers and non-normal models (e.g. Student-t)
- Dynamic models, e.g. AR(F)IMA and DLMs
- Models with endogenous regressors (IV models)
- Models for longitudinal data with individual effects
- Models for spatial data (Spatial AR models)
- Duration models

## Endogeneity

Endogeneity occurs if one or more of the covariates is correlated with the error term in the equation corresponding to (5). In particular, consider the following extension of the model in (5):

$$y = \alpha \iota + x \gamma + Z_j \beta_j + \varepsilon \tag{9}$$

$$x = W\delta + \nu, \tag{10}$$

where x is an endogenous regressor and W is a set of instruments, independent of  $\varepsilon$ . The error terms are iid:

$$(\varepsilon_i, \nu_i)' \sim N(0, \Sigma),$$
 (11)

with  $\Sigma=(\sigma_{ij})$  a  $2\times 2$  covariance matrix. Whenever  $\sigma_{12}\neq 0$  this introduces a bias in the OLS estimator of  $\gamma$  and a standard classical approach is the use of 2SLS estimators instead.

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Karl and Lenkoski (2012) propose IVBMA, which uses conditional Bayes factors to account for model uncertainty within a Gibbs algorithm. Their algorithm hinges on certain restrictions (e.g. joint Normality and conditionally conjugate priors), but it is exact, efficient and is implemented in an R-package.

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In previous literature some (influential) studies found evidence that property rights are a strong driver for growth, but without considering many alternative models. Similarly, others concluded that trade variables were key drivers (without controlling for the effect of institutions). Rodrik et al. (2004) (RST) provide a "horse race" among alternative theories that propose candidate instruments and regressors, but don't use BMA (they just compare a limited set of models) and conclude only institutions matter

Lenkoski et al. (2014). Econometric Reviews

Consider e.g. Rule of Law and Integration (Openness):

	Rule of Law			Integration		
Models	PIP	mean	sd	PIP	mean	sd
RST core	1.00	1.28	0.18	0.20	0.11	0.26
limited ${\cal M}$	1.00	0.95	0.13	0.07	0.07	0.14
full ${\cal M}$	0.96	0.80	0.32	0.85	0.93	0.38

Table: Some BMA (2nd stage) results with different sets of possible covariates (PIP is posterior inclusion probability)

Divergence of results (between 2SLS and BMA) grows as we allow for more uncertainty (bigger model spaces). Integration becomes important driver and all three theories are supported in the BMA results using all available variables.

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Błazejowski and Kwiatkowski (2015, 2018) present packages that implement BMA (including jointness measures) and BACE (amounts to BMA with a particular implicit choice of g) in gretl

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- Priors matter for BMA and it is crucial to be aware of this
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- BMA can become a key methodology in many areas of application, and can contribute to constructive communication by better understanding the reasons for differences in empirical findings

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